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## 基于临界 Galton-Watson 过程的 随机游动的大偏差

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**摘要:** 针对一族独立同分布的随机变量  $\{X_k\}$  的和  $S_{Z_n} = \sum_{k=1}^{Z_n} X_k$  ( $Z_n$  为临界 Galton-Watson 过程的第  $n$  代个体数), 利用随机游动和概率论的知识研究了  $R_n := S_{Z_n}/Z_n$  的渐近性质以及在  $\{Z_n > 0\}$  条件下的  $S_{Z_n}$  的大偏差. 研究表明,  $R_n$  的规范偏差概率有非退化的极限, 并且其大偏差规范化后收敛到正常数.

**关键词:** 大偏差概率; 临界 Galton-Watson 过程; 条件概率; Fuk-Nagaev 不等式

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## Large deviation of random walk indexed by critical Galton-Watson process

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**Abstract:** Consider the sum  $S_{Z_n} = \sum_{k=1}^{Z_n} X_k$  of a family of independent and identically distributed random variables  $\{X_k\}$ , where  $Z_n$  is the number of individuals of a critical Galton-Watson process in the generation  $n$ . The asymptotic properties of the ratio  $R_n := S_{Z_n}/Z_n$  and the large deviation of  $S_{Z_n}$  under  $\{Z_n > 0\}$  are studied by using the properties of random walks and branching processes. The proofs are given in detail that the normal deviation probability of  $R_n$  has nondegenerate limit and the large deviation probabilities of  $R_n$  after normalization converge to a positive constant.

**Keywords:** large deviation probability; critical Galton-Watson process; conditional probability; Fuk-Nagaev inequality

### 0 引言

对后代均值  $m$  的估计是研究 Galton-Watson 过程的一个重要内容. 令  $Z = (Z_n)_{n \geq 0}$  是后代分布为  $\{p_k; k \geq 0\}$  的 Galton-Watson 过程, 且假设  $Z$  是临界的, 即  $m := \sum_{k=1}^{\infty} kp_k = 1$ . 在文献[1]中, Nagaev 提出了著名的 Lotka-Nagaev 估计, 即用  $\frac{Z_{n+1}}{Z_n}$  来估计  $m$ . 如果  $\xi := (\text{Var } Z_1)^{1/2} \in (0, \infty)$ , 那么对所有  $x \in \mathbf{R}$ , 有

$$\lim_{n \uparrow \infty} \mathbf{P}\left(m^{n/2} \left(\frac{Z_{n+1}}{Z_n} - m\right) < x; Z_n > 0\right) = \int_0^{\infty} \Phi\left(\frac{xu^{1/2}}{\xi}\right) \omega(u) du. \quad (1)$$

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其中在  $\{W > 0\}$  条件下,  $\omega$  表示几乎处处收敛的变量  $W := \lim_{n \uparrow \infty} Z_n/m^n$  的连续密度函数,  $\Phi$  是标准正态分布函数, 可表示为  $\Phi(y) := \frac{1}{\sqrt{2\pi}} \int_{-\infty}^y e^{-\frac{z^2}{2}} dz, y \in \mathbf{R}$ .

Galton-Watson 过程是一类特殊的 Markov 链, 它被广泛应用在算法、数据结构、生物学、物理学等领域. 有学者研究发现, 在生物科学研究中比值  $\frac{Z_{n+1}}{Z_n}$  可以对定量的聚合酶链反应实验的扩散速率、初始值大小等重要参数进行估计. 例如: 文献[1] 对齐次分支过程的后代均值进行了估计. 文献[2] 通过定义灭绝概率  $q$  的一个估计序列  $p_j (j=0, 1, 2, \dots)$ , 证明了在  $\alpha = f'(1-q) \geq 1$  时估计序列  $p_j (j=0, 1, 2, \dots)$  具有一致性和渐近正态性. 文献[3] 在  $\theta$ -迁移过程为上临界 Galton-Watson 过程的假设下, 得到了具有迁移情况的上临界 Galton-Watson 过程的渐近性以及  $m$  和  $\sigma^2$  的估计结果. 文献[4] 对上临界的 Galton-Watson 过程的后代参数进行了估计, 同时考虑了大偏差和下偏差对其估计值的影响. 文献[5] 对突变的聚合酶链反应模型中的类似  $S_{Z_n}$  的和  $\sum_{x=1}^{S_0} u^{(x)}$  进行了研究, 结果显示突变的聚合酶链反应模型具有很强的平均效应, 即使对非常小的初始种群也是如此. 受文献[1] 和文献[2] 的启发, 本文考虑下述模型: 令  $X = (X_n)_{n \geq 1}$  表示独立同分布的(实值)随机变量族, 且与  $Z$  独立; 每一个  $X_n$  的均值都为 0, 方差都有限. 令  $S_n := X_1 + X_2 + \dots + X_n, n \geq 0$ , 则在  $\{Z_n > 0\}$  条件下, 随机变量  $R_n := \frac{S_{Z_n}}{Z_n}$  是有意义的.

依照文献[5] 中的方法, 很容易验证  $R_n$  的正态大偏差概率为

$$\lim_{n \uparrow \infty} P(m^{n/2} R_n < x) = \int_0^x \Phi\left(\frac{xu^{1/2}}{\sigma}\right) \omega(u) du, x \in \mathbf{R}, \quad (2)$$

其中  $\sigma := [E(X_1^2)]^{1/2}$ . 令  $\epsilon_n > 0$ , 考虑  $P(R_n \geq \epsilon_n)$  的相关问题. 当  $\epsilon_n m^{n/2} \rightarrow \infty$  时, 则由式(2) 知  $R_n$  具有简单的大偏差概率,  $\lim_{n \uparrow \infty} P(R_n \geq \epsilon_n) = 0$ .

本文在更一般的情况下(当  $n \uparrow \infty$  时,  $\epsilon_n \rightarrow 0$ ) 利用  $Z$  的下偏差概率的结果来研究  $R_n \geq \epsilon_n$  的概率收敛速率(大偏差), 并且在文献[4] 研究结果的基础上考虑在临界 Galton-Watson 过程下  $P(n^{1/2} R_n < x | Z_n > 0)$  的概率极限和在  $\epsilon_n \rightarrow 0, \epsilon_n^2 n \rightarrow \infty$  条件下  $\epsilon_n^2 n P(R_n \geq \epsilon_n | Z_n > 0)$  的收敛速率, 并给出其证明. 本文所得结论可为估计突变聚合酶链反应实验中的初始粒子、扩散速率等参数提供理论基础.

## 1 相关假设及引理

### 1.1 基本假设

假设  $E(X_1^2) < \infty, E(Z_1 \log Z_1) < \infty$ , 并考虑  $\epsilon_n > 0, \epsilon_n \rightarrow 0$  和当  $n \uparrow \infty$  时,  $\epsilon_n^2 n \rightarrow \infty$  的情况.

### 1.2 相关结果

引理 1<sup>[4]</sup> (Fuk-Nagaev 不等式) 对  $k \geq 1, \epsilon_n > 0, n \geq 1, r > 1$ , 有

$$P(S_k \geq \epsilon_n k) \leq k P(X_1 \geq r^{-1} \epsilon_n k) + (e r \sigma^2)^r \epsilon_n^{-2r} k^{-r}.$$

引理 2<sup>[6]</sup> (基本引理) 如果  $m = E Z_1 = 1, \xi^2 = \text{Var } Z_1 < \infty$ , 那么在  $0 \leq t < 1$  上一致有

$$\lim_{n \rightarrow \infty} \frac{1}{n} \left[ \frac{1}{1 - f_n(t)} - \frac{1}{1 - t} \right] = \frac{\xi^2}{2},$$

其中  $f_n(t)$  为矩母函数  $f(t)$  的  $n$  次迭代.

注 1 如果  $m = 1, \xi^2 < \infty$ , 那么当  $n \rightarrow \infty$  时有  $1 - f_n(0) = P\{Z_n > 0\} \sim \frac{2}{\xi^2 n} \triangleq c n^{-1} (c > 0 \text{ 为常数})$ .

定理 1<sup>[6]</sup> 如果  $m = 1, \xi^2 < \infty$ , 那么有  $\lim_{n \uparrow \infty} P\left(\frac{Z_n}{n} > z | Z_n > 0\right) = \exp\left\{-\frac{2z}{\xi^2}\right\}, z > 0$ .

**定理 2**<sup>[7]</sup> 令  $\xi^2 < \infty$ ,  $k$  和  $n$  的比值  $k/n$  以有界的方式趋于无穷, 那么有

$$\lim_{n \rightarrow \infty} \frac{\xi^4 n^2}{4d} \left(1 + \frac{2d}{\xi^2 n}\right)^{k+1} P(Z_n = kd) = 1,$$

其中  $d$  为  $\{p_k; k \geq 0\}$  的最大公约数.

**引理 3** 对所有  $n, k \geq 1$ , 存在一个常数  $C_1$ , 使得  $P(Z_n = k | Z_n > 0) \leq \frac{C_1}{n}$ .

**证明** 令  $t=0$ , 则根据引理 2 有  $\lim_{n \rightarrow \infty} \frac{1}{n} \left[ \frac{1}{1 - f_n(0)} - 1 \right] = \frac{\xi^2}{2}$ , 即  $\lim_{n \rightarrow \infty} \frac{1}{nP\{Z_n > 0\}} = \frac{\xi^2}{2}$ . 根据文献 [7] 中的性质知, 若  $\xi^2 < \infty$ , 则存在常数  $c_1$  使得  $\forall k, n \geq 1, \sup_{n, k \geq 1} n^2 P(Z_n = k) \leq c_1$ , 因此有  $P(Z_n = k) \leq$

$\frac{C_2}{n^2}$ . 所以当  $n$  足够大时有  $P(Z_n = k | Z_n > 0) = \frac{nP(Z_n = k)}{nP(Z_n > 0)} \leq \frac{n \frac{C_2}{n^2}}{nP(Z_n > 0)} \leq \frac{C_2 \xi^2}{n} \triangleq \frac{C_1}{n}$ . 引理证毕.

## 2 主要结论及其证明

**定理 3** 假设  $Z$  是临界 Galton-Watson 过程,  $m=1$ ,  $\xi^2 = \text{Var } Z_1 \in (0, \infty)$ , 则根据式 (2) 有

$$\lim_{n \uparrow \infty} P(n^{1/2} R_n < x | Z_n > 0) = \frac{2}{\xi^2} \int_0^\infty \Phi\left(\frac{xu^{1/2}}{\sigma}\right) e^{-\frac{2x}{\xi^2} u} du. \quad (3)$$

**证明** 根据全概率公式知, 对  $\forall \delta \in (0, 1)$  和  $1 < A < \infty$  有

$$\begin{aligned} P(n^{1/2} R_n < x | Z_n > 0) &= \sum_{k=1}^{\infty} P(n^{1/2} S_k < xk) P(Z_n = k | Z_n > 0) = \\ &= \sum_{k < \delta n} P(n^{1/2} S_k < xk) P(Z_n = k | Z_n > 0) + \sum_{\delta n \leq k \leq An} P(n^{1/2} S_k < xk) P(Z_n = k | Z_n > 0) + \\ &= \sum_{k > An} P(n^{1/2} S_k < xk) P(Z_n = k | Z_n > 0) \triangleq a + b + c. \end{aligned}$$

以下分 3 部分证明定理 3.

(a) 根据引理 3, 对  $\forall k, n \geq 1$ , 有

$$\sum_{k < \delta n} P(n^{1/2} S_k < xk) P(Z_n = k | Z_n > 0) \leq \sum_{k < \delta n} P(Z_n = k | Z_n > 0) \leq \sum_{k < \delta n} \frac{C_3}{n} = C_3 \delta. \quad (4)$$

(b) 当  $\delta n \leq k \leq An$ , 且当  $n$  足够大时, 令  $d=1$ , 则根据定理 2 可知  $P(Z_n = k) = (1 + o(1)) \cdot \frac{4}{\xi^4 n^2} \left(1 + \frac{2}{\xi^2 n}\right)^{-(k+1)}$ . 再由引理 2 可得  $\lim_{n \rightarrow \infty} \frac{1}{nP(Z_n > 0)} = \frac{\xi^2}{2}$ . 所以, 当  $\delta n \leq k \leq An$ , 且  $n$  足够大时有

$P(Z_n = k | Z_n > 0) = (1 + o(1)) \frac{1}{n} \frac{2}{\xi^2} e^{-\frac{2}{\xi^2} (\frac{k}{n})}$ . 进而可得

$$\begin{aligned} \sum_{\delta n \leq k \leq An} P(n^{1/2} S_k < xk) P(Z_n = k | Z_n > 0) &= (1 + o(1)) \sum_{\delta n \leq k \leq An} P\left(\frac{S_k}{\sigma \sqrt{k}} < \frac{x \sqrt{k}}{\sigma \sqrt{n}}\right) \frac{1}{n} \frac{2}{\xi^2} e^{-\frac{2}{\xi^2} (\frac{k}{n})} = \\ &= (1 + o(1)) \sum_{\delta n \leq k \leq An} \Phi\left(\frac{x \sqrt{k}}{\sigma \sqrt{n}}\right) \frac{1}{n} \frac{2}{\xi^2} e^{-\frac{2}{\xi^2} (\frac{k}{n})} = (1 + o(1)) \int_{\delta}^A \Phi\left(\frac{x \sqrt{u}}{\sigma}\right) \frac{2}{\xi^2} e^{-\frac{2}{\xi^2} u} du = \\ &= (1 + o(1)) \frac{2}{\xi^2} \int_{\delta}^A \Phi\left(\frac{x \sqrt{u}}{\sigma}\right) e^{-\frac{2}{\xi^2} u} du. \end{aligned} \quad (5)$$

(c)  $\sum_{k > An} P(n^{1/2} S_k < xk) P(Z_n = k | Z_n > 0) \leq \sum_{k > An} P(Z_n = k | Z_n > 0) =$

$$1 - \sum_{k < An} P(Z_n = k | Z_n > 0). \quad (6)$$

因  $\sum_{k < An} P(Z_n = k | Z_n > 0) = 1$ , 所以当  $n$  足够大时, 式 (6) 等于  $o(1)$ . 当  $A \rightarrow \infty$  时, 式 (6) 的极限趋于 0.

当  $\delta \rightarrow 0^+$  时, 式(4) 的极限趋于 0. 当  $\delta \rightarrow 0^+$ ,  $A \rightarrow \infty$  时,

$$\sum_{\delta n \leq k \leq An} P(n^{1/2} S_k < xk) P(Z_n = k | Z_n > 0) \rightarrow \frac{2}{\xi^2} \int_{\delta}^A \Phi\left(\frac{x\sqrt{u}}{\sigma}\right) e^{-\frac{2}{\xi^2}u} du.$$

综上所述, 定理 3 证毕.

**定理 4** 假设  $\epsilon_n \rightarrow 0$ ,  $\epsilon_n^2 n \rightarrow \infty$ , 则  $\lim_{n \uparrow \infty} \epsilon_n^2 n P(R_n \geq \epsilon_n | Z_n > 0) = \frac{\sigma^2}{\xi^2}$ .

**证明** 根据全概率公式知, 对  $\forall \delta \in (0, 1)$  和  $1 < A < \infty$  有

$$\begin{aligned} \epsilon_n^2 n P(R_n \geq \epsilon_n | Z_n > 0) &= \epsilon_n^2 n \sum_{k=1}^{\infty} P\left(\sum_{i=1}^k X_i \geq \epsilon_n k, Z_n = k | Z_n > 0\right) = \\ &= \epsilon_n^2 n \sum_{k=1}^{\infty} P\left(\sum_{i=1}^k X_i \geq \epsilon_n k | Z_n = k\right) P(Z_n = k | Z_n > 0) = \\ &= \epsilon_n^2 n \sum_{k < \delta/\epsilon_n^2} P\left(\sum_{i=1}^k X_i \geq \epsilon_n k | Z_n = k\right) P(Z_n = k | Z_n > 0) + \\ &= \epsilon_n^2 n \sum_{\delta/\epsilon_n^2 < k < A/\epsilon_n^2} P\left(\sum_{i=1}^k X_i \geq \epsilon_n k | Z_n = k\right) P(Z_n = k | Z_n > 0) + \\ &= \epsilon_n^2 n \sum_{k > A/\epsilon_n^2} P\left(\sum_{i=1}^k X_i \geq \epsilon_n k | Z_n = k\right) P(Z_n = k | Z_n > 0) \triangleq \text{I} + \text{II} + \text{III}. \end{aligned}$$

下面分 3 部分证明定理 4.

$$\begin{aligned} \text{(I)} \quad \lim_{n \uparrow \infty} \epsilon_n^2 n \sum_{k < \delta/\epsilon_n^2} P\left(\sum_{i=1}^k X_i \geq \epsilon_n k | Z_n = k\right) P(Z_n = k | Z_n > 0) &\leq \\ \lim_{n \uparrow \infty} \epsilon_n^2 n \sum_{k < \delta/\epsilon_n^2} P(Z_n = k | Z_n > 0) &= \lim_{n \uparrow \infty} (1 + o(1)) \epsilon_n^2 n \sum_{k < \delta/\epsilon_n^2} P(Z_n = k) \frac{\xi^2 n}{2} = \\ \lim_{n \uparrow \infty} (1 + o(1)) \frac{\xi^2 n^2 \epsilon_n^2}{2} \sum_{k < \delta/\epsilon_n^2} \frac{4}{\xi^4 n^2} \left(1 + \frac{2}{\xi^2 n}\right)^{-(k+1)} &= (1 + o(1)) \frac{2\epsilon_n^2}{\xi^2} \sum_{k < \delta/\epsilon_n^2} (e^{2/\xi^2})^{-\frac{k+1}{n}} \leq \\ (1 + o(1)) \frac{2\epsilon_n^2}{\xi^2} \sum_{k < \delta/\epsilon_n^2} C_4^{-\frac{k}{n}} &\leq (1 + o(1)) \frac{2\epsilon_n^2}{\xi^2} \frac{\delta}{\epsilon_n^2} = \frac{2\delta}{\xi^2} \rightarrow 0. \\ \text{(II)} \quad \lim_{n \uparrow \infty} \epsilon_n^2 n \sum_{\delta/\epsilon_n^2 < k < A/\epsilon_n^2} P\left(\sum_{i=1}^k X_i \geq \epsilon_n k | Z_n = k\right) P(Z_n = k | Z_n > 0) &= \\ \epsilon_n^2 n \sum_{\delta/\epsilon_n^2 < k < A/\epsilon_n^2} \lim_{n \uparrow \infty} P\left(\frac{S_k}{\sigma\sqrt{k}} < \frac{\epsilon_n k}{\sigma\sqrt{k}}\right) P(Z_n = k | Z_n > 0) &= \\ \epsilon_n^2 n \sum_{\delta/\epsilon_n^2 < k < A/\epsilon_n^2} \bar{\Phi}\left(\frac{\epsilon_n \sqrt{k}}{\sigma}\right) P(Z_n = k | Z_n > 0) &= \\ (1 + o(1)) \epsilon_n^2 n \sum_{\delta/\epsilon_n^2 < k < A/\epsilon_n^2} \bar{\Phi}\left(\frac{\epsilon_n \sqrt{k}}{\sigma}\right) \frac{1}{n} \frac{2}{\xi^2} e^{-\frac{2}{\xi^2}(\frac{k}{n})} &= (1 + o(1)) \frac{2}{\xi^2} \int_{\delta/\epsilon_n^2}^{A/\epsilon_n^2} \epsilon_n^2 \bar{\Phi}\left(\frac{\epsilon_n \sqrt{u}}{\sigma}\right) du = \\ (1 + o(1)) \frac{2}{\xi^2} \int_{\delta}^A \bar{\Phi}\left(\frac{\sqrt{x}}{\sigma}\right) dx &\leq \frac{2}{\xi^2} \int_0^{\infty} \bar{\Phi}\left(\frac{\sqrt{x}}{\sigma}\right) dx = \frac{2}{\xi^2} \frac{1}{2} \sigma^2 = \frac{\sigma^2}{\xi^2}, \end{aligned}$$

其中  $\bar{\Phi}\left(\frac{\sqrt{x}}{\sigma}\right) = 1 - \Phi\left(\frac{\sqrt{x}}{\sigma}\right)$ .

(III) 根据引理 1, 令  $r=2$ , 则有  $P(S_k \geq \epsilon_n k) \leq kP\left(X_1 \geq \frac{\epsilon_n k}{2}\right) + (2e\sigma^2)^2 \epsilon_n^4 k^{-2}$ . 再结合引理 3 可得

$$\epsilon_n^2 n \sum_{k > A/\epsilon_n^2} P\left(\sum_{i=1}^k X_i \geq \epsilon_n k | Z_n = k\right) P(Z_n = k | Z_n > 0) \leq \frac{C_5}{n} \epsilon_n^2 n \sum_{k > A/\epsilon_n^2} P(Z_n = k | Z_n > 0) \leq$$

$$C_5 \epsilon_n^2 \sum_{k > A/\epsilon_n^2} \left[ k P\left(X_1 \geq \frac{\epsilon_n k}{2}\right) + (2e\sigma^2)^2 \epsilon_n^{-4} k^{-2} \right] \leq \\ C_5 \epsilon_n^2 \sum_{k > A/\epsilon_n^2} \left[ k P\left(X_1 \geq \frac{\epsilon_n k}{2}\right) \right] + C_6 \epsilon_n^2 \sum_{k > A/\epsilon_n^2} \epsilon_n^{-4} k^{-2} \triangleq D + E.$$

根据文献[4]中的式(76),令  $\alpha = 1$ , 则有  $\sum_{k > A/\epsilon_n^2} \epsilon_n^{-4} k^{-2} \leq \frac{C_7}{A} \epsilon_n^{-2}$ ,  $n > 0$ ,  $\epsilon_n > 0$ ,  $A \geq 1$ . 从而 D 部分变为

$$C_6 \epsilon_n^2 \sum_{k > A/\epsilon_n^2} \epsilon_n^{-4} k^{-2} \leq C_6 \epsilon_n^2 \frac{C_7}{A} \epsilon_n^{-2} \leq \frac{C_8}{A} \rightarrow 0, \quad 1 < A \rightarrow \infty.$$

对  $k \geq 2$ , 根据  $\frac{k}{2} \leq k-1$  可得  $(k-1) P\left(X_1 \geq \frac{\epsilon_n k}{2}\right) \leq \int_{k-1}^k (x+1) P\left(X_1 \geq \frac{\epsilon_n x}{2}\right) dx$  是连续的, 从而有

$$C_5 \epsilon_n^2 \sum_{k > A/\epsilon_n^2} \left[ k P\left(X_1 \geq \frac{\epsilon_n k}{2}\right) \right] \leq C_5 \epsilon_n^2 \sum_{k > A/\epsilon_n^2} \int_{k-1}^k (x+2) P\left(X_1 \geq \frac{\epsilon_n x}{2}\right) dx = \\ C_5 \epsilon_n^2 \int_{A/\epsilon_n^2-1}^{\infty} (x+2) P\left(X_1 \geq \frac{\epsilon_n x}{2}\right) dx = \\ C_5 \epsilon_n^2 \int_{A/\epsilon_n^2-1}^{\infty} x P\left(X_1 \geq \frac{\epsilon_n x}{2}\right) dx + C_5 \epsilon_n^2 \int_{A/\epsilon_n^2-1}^{\infty} 2 P\left(X_1 \geq \frac{\epsilon_n x}{2}\right) dx = \\ C_5 \epsilon_n^2 \int_{A/\epsilon_n^2-1}^{\infty} x P\left(X_1 \geq \frac{\epsilon_n x}{2}\right) dx + (1+o(1)) C_5 \epsilon_n^2 \int_{A/\epsilon_n^2-1}^{\infty} 2 \frac{2}{\epsilon_n^2 x^2} dx = \\ C_5 \epsilon_n^2 \int_{A/\epsilon_n^2-1}^{\infty} x P\left(X_1 \geq \frac{\epsilon_n x}{2}\right) dx + o(1).$$

在上述式子中令  $v = \frac{x\epsilon_n}{2}$ , 则可得

$$C_5 \epsilon_n^2 \sum_{k > A/\epsilon_n^2} \left[ k P\left(X_1 \geq \frac{\epsilon_n k}{2}\right) \right] \leq (1+o(1)) C_5 \epsilon_n^2 \int_{\frac{A}{2\epsilon_n} - \frac{\epsilon_n}{2}}^{\infty} \frac{\epsilon_n}{2} v \frac{2}{\epsilon_n} P(X_1 \geq v) dv = \\ (1+o(1)) C_5 \epsilon_n^2 \int_{\frac{A}{2\epsilon_n} - \frac{\epsilon_n}{2}}^{\infty} v P(X_1 \geq v) dv.$$

又由于本文假设二阶矩有限, 因此当  $\frac{A}{2\epsilon_n} - \frac{\epsilon_n}{2} \rightarrow \infty$  时, 上述式子等于  $o(1)$ . 进而有

$$\epsilon_n^2 n \sum_{k > A/\epsilon_n^2} P\left(\sum_{i=1}^k X_i \geq \epsilon_n k \mid Z_n = k\right) P(Z_n = k \mid Z_n > 0) = o(1).$$

综上所述,  $\lim_{n \uparrow \infty} \epsilon_n^2 n P(R_n \geq \epsilon_n \mid Z_n > 0) = 0 + o(1) + \frac{\sigma^2}{\xi^2} = \frac{\sigma^2}{\xi^2}$ . 定理 4 证毕.

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